

UNDER REACTION

Regression Models for Repeated Measurements

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Response to a Query, Aitkin (*Biometrics* **37**, 831–832, December 1981) uses ANOVA methodology to analyze data from a trial in which rectal temperatures were measured on a group of ten subjects, under four combinations of ambient temperatures with six equally-spaced observations taken for each temperature. I feel that the assumptions needed for the 'standard' ANOVA are almost certainly not satisfied for this comparable data sets. To simplify the discussion, I assume that the ambient temperatures were applied in random order. Under this condition, the tests for temperature and for the linear components of the time effect and interaction, do have the stated F distribution if modified slightly, whereas the other tests do not have an F distribution.

It suggests a partition of the residual term to investigate the possibility that the variances between the subjects. The partition should be done even if the slope is not of interest, since it permits analysis of a more general model which includes Subject \times Time and Subject \times Temperature interactions (Neter and Wasserman, 1974, p. 732). In addition, as indicated below, it allows evaluation of the temperature effect, even if serial correlation due to time is present.

Although the necessary and sufficient condition for validity of the model does not require equal variances and covariances, it does imply that the covariances are equal if the variances are equal (Huynh and Feldt, 1970). Furthermore, Huynh (1978) states that 'it is difficult to conceptualize a situation which would give rise to a Huynh–Feldt situation'. I suggest that for all practical purposes the required assumption should be that of equal variances and equal covariances. This assumption is probably not satisfied for the data of time in this data set, since the correlation between the 20- and 40-minute observations in rectal temperature is probably not the same as between 20 minutes and two hours. Since the ambient temperatures were applied in random order, it would be reasonable to assume that the correlations between temperatures are equal and thus MS {Temperature}/ MS {Subject \times Temperature} has an F distribution (see also Mendoza, 1976). Since the test for trend in the time effect is based on a linear regression of the within-individual observations, MS {Time(L)}/ MS {Subject \times Time(L)} has an F distribution.

Huynh (1978) cites various approaches for giving approximate critical values for the F distribution where the Huynh–Feldt assumptions do not hold. If one can postulate that the correlation between observations u time units apart is ρ^u , conservative critical values are obtained by multiplying the nominal degrees of freedom by $5(t+1)/(2t^2+7)$ (Huynh and Fleiss, 1979).

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Murray Aitkin's Response to a Query on the use of regression in the analysis of experiments with repeated measurements (*Biometrics* **37**, 831-832, December 1981) requires some comment.

As a specific example the Query instanced 'an experiment in which rectal temperatures were measured on a group of 10 subjects at 20-minute intervals over exposure periods of two hours, at four different ambient temperatures'. Aitkin states that experiments of this kind are called 'split-plot' experiments and refers to Cochran and Cox (1953) for a detailed discussion of their analysis. This is incorrect terminology. When different treatments are applied to the subplots constituting a whole plot experiment, the experiment is properly randomized these treatments are randomly assigned to the subplots. In an experiment with repeated measurements ('successive' might be a better term) the measurements have a temporal sequence, with the consequence that measurements on the same subject separated by a small time interval will in general be highly correlated than those more widely separated. Cochran and Cox do not discuss the analysis of experiments with repeated measurements.

In the analysis of agronomic experiments it has long been recognised that when measurements of two or more different quantities are made plot by plot, separate analyses of variance of the plot-by-plot values of each quantity give valid estimates of error and tests of significance for the treatment means for each quantity, but that if a function of these quantities is also of interest, separate values of the function must be calculated for each plot to provide a basis for a further valid analysis. Experiments on sugar beet are a simple example. In these a sample of roots is taken from each plot to determine the plot-by-plot values of the sugar percentage. These, multiplied by the plot-by-plot yields of roots, give estimates of the plot-by-plot yields of sugar; it is these values that must be used for an analysis of variance relating to the effects of treatments on the sugar yield.

The same principle applies to repeated measurements. If, for example, the effects of treatments on the rates of change are of interest, linear regressions can be fitted to the measurements on each experimental plot. An analysis of variance of the values of the resultant regression coefficients can then be made. The subject is well discussed by Cochran and Walters (1976), with particular reference to agronomic experiments harvested on a number of successive occasions, such as those on fruit and vegetables.

The present experiment is analogous to an ordinary randomized block experiment with four treatments (temperatures) in ten blocks (subjects) with repeated measurements.

plot experiment. The only difference is that in an agronomic experiment there are plots in each block which are treated simultaneously, whereas in the present experiment the same subject is subjected successively to the four temperatures. If the temperatures are applied in random order to each subject individually an ordinary randomized block analysis of any function of the repeated measurements is valid.

To simplify the experiment, all the subjects are placed simultaneously in a single large chamber, this randomization requirement cannot be met. The same analysis is appropriate, but any residual effects of previous treatments will be confounded with the effects of the current treatment, instead of being included in experimental error. If a long interval is allowed to elapse between treatment periods such residual effects are to be unimportant; in other types of experiment, in which they are of importance, crossover designs, in which changes of treatment are incorporated in the design, with or without repeated measurements on the same treatment, are available. The term 'repeated measurements' must not be used to denote a crossover design.

In the present experiment the six values for any one subject exposed to a particular temperature can be fitted exactly by a fifth-degree polynomial. This can be done by its mean and five orthogonal polynomials of degree one to five; that of degree one is the best-fitting linear regression, that of degree two the improvement in fitting a quadratic term, etc. A separate analysis of variance (or a single analysis of variance in sections) can then be made on each of these functions. Because they are orthogonal linear functions, the total sum of squares for the six analyses will equal the sum of squares of deviations for all 240 observations, though orthogonality is not necessary for separate analyses to be valid.

Table below gives the degrees of freedom for these analyses. To save space, and for comparability with Aitkin's table, the analyses for quadratic and higher terms are lumped together. Also for comparability I have accepted Aitkin's supposition that only six values are observed in each test, though a wise experimenter would, I am sure, have included an additional observation at the start of the test.

	Means over all times	Regressions on time	
		Linear	Nonlinear
Mean	[cfm]	1	4
Subjects	9	9*	36*
Ambient temperatures	3	3	12
Subjects × Temperatures	27*	27*	108*

Items marked with an asterisk in my table are all included in the 'Residual within item' (207 df) in Aitkin's table, and by implication the items not so marked can be tested for significance against this residual term. In fact, as explained above, items can be properly tested against the 27 df for Subjects × Temperatures in the same column (the column being split into four columns for quadratic, cubic, etc. effects). Further, if the subjects react differently to the different ambient temperatures, and if linear regressions express these reactions 'adequately', the Subjects × Temperatures mean squares in the first two columns of my table, and the Subjects mean squares of the second column are likely to be substantially larger than the corresponding mean squares for the first column. This may well result in the values of the mean squares for the 4 df and 12 df nonlinear components of the means and temperatures not attaining significance,

even at a low level, when tested against Aitkin's residual, thus giving a false impression of the adequacy of the linear regressions. More importantly, the significance of the average effects of the different ambient temperatures will be greatly exaggerated, and the corresponding standard errors will be similarly underestimated.

Aitkin does point out that to investigate the possibility that the slope varies between subjects his residual term could, 'if necessary', be subdivided into the asterisked term in my table, but does not indicate to the reader that the Subjects linear term (9 df) should be tested against the Subjects \times Temperatures linear term (27 df), not the residual term (108 df). It is of course true that the Subjects \times Temperatures terms may contain genuine repeatable subject differences as well as experimental error: subjects may react differently to different temperatures because of physiological differences, such as those due to age or sex. An indication of such differences might be obtained by extracting the linear component of Subjects \times Temperature, or subdivision by, or covariance on, temperature and sex.

Although an analysis of the above type, using orthogonal polynomials, is conceptually and computationally simple, it may not provide the best approach in specific problems. In the present example a preliminary plot of the results (which should always be made) may well indicate that changes in rectal temperatures over time are of the sigmoid type, asymptotic to an initial level (which may differ slightly from subject to subject), and asymptotic to a changed level dependent on the change in ambient temperature and on the subject's physiological reactions. If this is so, the main quantities of interest will probably be (i) the changes in level and (ii) the speed with which the new level is attained.

Evidence of sigmoid-type curves will indeed be provided by the existence of significant quadratic and cubic components when orthogonal polynomials are fitted as above, but the differences of the maxima and minima of third degree polynomials are unlikely to give satisfactory estimates of the changes in level. If a suitable computer program such as the Rothamsted Maximum Likelihood Program (MLP) is available the ideal solution might be to fit a Gompertz curve (which does not assume equal rates of convergence to the horizontal asymptotes) to each sequence of observations. An alternative and very simple procedure is to estimate the change of level for each subject from the differences of the mean of the last few observations (e.g. the last two) from the initial observation. An indication of differences in rates of change between subjects will then be obtained from the ratios of the means for the subjects over all times to the corresponding changes of level. When the sets of quantities have been calculated a separate analysis of variance can be made for each set.

It may be objected that this latter procedure does not extract the full amount of information available in the data. The information which is extracted however, is readily intelligible, and the particular contrasts selected can be varied in the light of the form and regularity of the observed curves. That a wise choice is capable of recovering most of the available information is illustrated at an elementary level by the estimation of a regression coefficient from the difference of the means of the first and last two or three observations of a short series of equally-spaced observations. If the regression is truly linear and the errors are independent, the loss of information for series of seven, nine and 11 observations and means of two observations are 11%, 18% and 26%; for means of three they are 14%, 10% and 13%. Means of two are therefore a good choice for seven observations and means of three for nine and 11 observations.

At the end of his paper Aitkin makes reference to various involved multivariate methods of analysis that have appeared in the literature. I will not comment on these methods except to say that I much doubt their relevance to real problems. His last paragraph

reads 'Other complications, like serial correlation between plots in a block, have to be taken into account by general maximum likelihood methods, though other simple approaches have been suggested' is particularly revealing. Correlations of many kinds, both between successive observations on the same units and between simultaneous observations on different units, are usually to be expected. An incorrect model will also introduce spurious correlations between residuals from basically independent observations, as for example when the true functions are second-degree polynomials but linear regression lines are fitted. It is, I believe, failure to appreciate these facts that is the main cause of Aitkin's over-reliance on the analysis of variance.

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RESPONSE

A brief response to a general question necessarily omitted qualifying details. Readers who have asked for more information will welcome the detailed comments by Wallenstein and Box (1976). A clear and careful exposition of the analysis of repeated measures designs is given by Box (1975, Chapter 7), who gives also the general multivariate analysis of variance when the covariance matrix of the repeated measures does not satisfy the condition for the validity of the standard analysis of variance. The necessary and sufficient condition is sphericity of the *differences* of pairs of the repeated measures: sphericity of the repeated measures themselves is sufficient, but not necessary. The general multivariate analysis of variance is valid, though not efficient, for particular patterns of covariance structure like compound symmetry.

Box also gives the analysis for individual polynomial contrasts of the repeated measures.

In computing, the statistical package GENSTAT is particularly useful for split-plot designs, as very simple declarations of the block, plot and treatment structure are sufficient to provide complete ANOVA breakdowns for generally balanced designs. For a compound correlation structure in repeated-measures designs, multivariate analysis of variance programs are required. These are available in most statistical packages.

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